

# Package ‘BAMBI’

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**Description** Fit (using Bayesian methods) and simulate mixtures of univariate and bivariate angular distributions. Chakraborty and Wong (2017) <arXiv:1708.07804> .

**License** GPL-3

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AIC.angmcmc	<i>AIC and BIC for angmcmc objects</i>
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---

**Description**

AIC and BIC for angmcmc objects

**Usage**

```
## S3 method for class 'angmcmc'
AIC(object, burnin = 1/3, thin = 1, k = 2, ...)
```

```
## S3 method for class 'angmcmc'
BIC(object, burnin = 1/3, thin = 1, ...)
```

**Arguments**

object	an angmcmc object.
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in [0, 1).
thin	positive integer. If thin = $n$ , only every $n$ -th realizations of the Markov chain is kept.
k	numeric, the <i>penalty</i> per parameter to be used; the default k = 2 is the classical AIC.
...	optionally more fitted model objects.

**Details**

Let  $\hat{L}$  be the maximum value of the likelihood function for the model,  $m$  be the number of estimated parameters in the model and  $n$  be the number of data points. Then AIC and BIC are defined as  $AIC = -2 \log \hat{L} + mk$  and  $BIC = -2 \log \hat{L} + m \log(n)$ .

$\hat{L}$  is estimated by the sample maximum obtained from the MCMC realizations.

**Value**

AIC computes the AIC and BIC computes BIC for angmcmc objects.

**Examples**

```
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)

AIC(fit.vmsin.20)
BIC(fit.vmsin.20)
```

---

bestmodel	<i>Extracting angmcmc object corresponding to the best fitted model in stepwise fits</i>
-----------	--

---

**Description**

Extracting angmcmc object corresponding to the best fitted model in stepwise fits

**Usage**

```
bestmodel(step_object)
```

**Arguments**

step\_object      stepwise fitted object (output of [fit\\_stepwise\\_univariate](#) or [fit\\_stepwise\\_bivariate](#)).

**Value**

Returns an `angmcmc` object corresponding to the the best fitted model in `step_object`.

**Examples**

```
# illustration only - more iterations needed for convergence
fit.vmsin.step.15 <- fit_stepwise_bivariate(tim8, "vmsin", start_ncomp = 3,
                                          max_ncomp = 5, n.iter = 15,
                                          ncores = 1)
fit.vmsin.best.15 <- bestmodel(fit.vmsin.step.15)
fit.vmsin.best.15
```

---

circ\_cor

*Sample circular correlation coefficients*

---

**Description**

Sample circular correlation coefficients

**Usage**

```
circ_cor(x, type = "js")
```

**Arguments**

<code>x</code>	two column matrix. NA values are not allowed.
<code>type</code>	type of the circular correlation. Must be one of "fl" or "js" or "kendall". See details.

**Details**

`circ_cor` calculates the (sample) circular correlation between the columns of `x`. Two parametric (the Jammalamadaka-Sarma (1988, equation 2.6) form "js", and the Fisher-Lee (1983, Section 3) form "fl") and two non-parametric (two versions of Kendall's tau) correlation coefficients are considered. The first version of Kendall's tau ("tau1") is based on equation 2.1 in Fisher and Lee (1982), whereas the second version ("tau2") is computed using equations 6.7-6.8 in Zhan et al (2017).

The cost-complexity for "js", "fl", "tau2" and "tau1" are  $O(n)$ ,  $O(n^2)$ ,  $O(n^2)$  and  $O(n^3)$  respectively, where  $n$  denotes the number of rows in `x`. As such, for large  $n$  evaluation of "tau1" will be slow.

## References

- Fisher, N. I. and Lee, A. J. (1982). Nonparametric measures of angular-angular association. *Biometrika*, 69(2), 315-321.
- Fisher, N. I. and Lee, A. J. (1983). A correlation coefficient for circular data. *Biometrika*, 70(2):327-332.
- Jammalamadaka, S. R. and Sarma, Y. (1988). A correlation coefficient for angular variables. *Statistical theory and data analysis II*, pages 349-364.
- Zhan, X., Ma, T., Liu, S., & Shimizu, K. (2017). On circular correlation for data on the torus. *Statistical Papers*, 1-21.

## Examples

```
# generate data from vmsin model
set.seed(1)
dat <- rvmsin(100, 2,3,-0.8,0,0)

# now calculate circular correlation(s) between the 2 columns of dat
circ_cor(dat, type="js")
circ_cor(dat, type="f1")
circ_cor(dat, type="tau1")
circ_cor(dat, type="tau2")
```

---

circ_varcor_model	<i>Analytic circular variances and correlations for bivariate angular models</i>
-------------------	--

---

## Description

Analytic circular variances and correlations for bivariate angular models

## Usage

```
circ_varcor_model(model = "vmsin", kappa1 = 1, kappa2 = 1, kappa3 = 0,
  mu1 = 0, mu2 = 0, nsim = 10000)
```

## Arguments

model	bivariate angular model. Must be one of "vmsin", "vmcos", or "wnorm2".
kappa1, kappa2, kappa3	concentration and covariance parameters. Recycled to the same size. kappa3 <sup>2</sup> must be < kappa1*kappa2 in the wnorm2 model (see <a href="#">rwnorm2</a> for a detailed parameterization of wnorm2).
mu1, mu2	mean parameters. Ignored as they do not play any role in the analytical formulas.
nsim	Monte Carlo sample size. Ignored if all of kappa1, kappa2 and abs(kappa3) are < 150 or if model = "wnorm2".

## Details

The function computes the analytic circular variances and correlations (both Jammalamadaka-Sarma (JS) and Fisher-Lee (FL) forms) for von Mises sine, von Mises cosine and bivariate wrapped normal distributions.

For `wnorm2`, expressions for the circular variances, JS and FL correlation coefficients can be found in Mardia and Jupp (2009), Jammalamadaka and Sarma (1988) and Fisher and Lee (1983) respectively. For `vmsin` and `vmcos` these expressions are provided in Chakraborty and Wong (2018).

Because the analytic expressions in `vmsin` and `vmcos` models involve infinite sums of product of Bessel functions, if any of `kappa1`, `kappa2` and `kappa3` is larger than or equal to 150, IID Monte Carlo with sample size `nsim` is used to approximate `rho_js` for numerical stability. From `rho_js`, `rho_fl` is computed using Corollary 2.2 in Chakraborty and Wong (2018), which makes cost-complexity for the `rho_fl` evaluation to be of order  $O(nsim)$  for `vmsin` and `vmcos` models. (In general, `rho_fl` evaluation is of order  $O(nsim^2)$ ).

## Value

Returns a list with elements `var1`, `var2` (circular variances for the first and second coordinates), `rho_fl` and `rho_js` (circular correlations). See details.

## References

- Fisher, N. I. and Lee, A. (1983). A correlation coefficient for circular data. *Biometrika*, 70(2):327-332.
- Jammalamadaka, S. R. and Sarma, Y. (1988). A correlation coefficient for angular variables. *Statistical theory and data analysis II*, pages 349-364.
- Mardia, K. and Jupp, P. (2009). *Directional Statistics*. Wiley Series in Probability and Statistics. Wiley.
- Chakraborty, S. and Wong, S. W.K. (2018). On the circular correlation coefficients for bivariate von Mises distributions on a torus. arXiv e-print.

## Examples

```
circ_varcor_model("vmsin", kappa1= 1, kappa2 = 2, kappa3 = 3)

# Monte Carlo approximation
set.seed(1)
dat <- rvmsin(1000, 1, 2, 3)
# sample circular variance
circ_var <- function(x)
  1 - mean(cos(x - atan2(mean(sin(x)), mean(cos(x)))) )
circ_var(dat[, 1])
circ_var(dat[, 2])
circ_cor(dat, "fl")
circ_cor(dat, "js")
```

---

contour.angmcmc	<i>Contour plot for angmcmc objects with bivariate data</i>
-----------------	---

---

**Description**

Contour plot for angmcmc objects with bivariate data

**Usage**

```
## S3 method for class 'angmcmc'
contour(x, fn = mean, show.data = TRUE, nlevels = 20,
        levels, burnin = 1/3, thin = 1, cex = 1, col = "red", ...)
```

**Arguments**

x	angular MCMC object (with bivariate data).
fn	function to evaluate on MCMC samples to estimate parameters. Defaults to mean, which computes the estimated posterior mean.
show.data	logical. Should the data points be added to the contour plot? Ignored if object is NOT supplied.
nlevels	number of contour levels desired <b>if</b> levels is not supplied; passed to the <a href="#">contour</a> function in graphics.
levels	numeric vector of levels at which to draw contour lines; passed to the <a href="#">contour</a> function in graphics.
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in [0, 1).
thin	positive integer. If thin = n, only every n-th realizations of the Markov chain is kept.
cex, col	graphical parameters passed to <a href="#">points</a> in graphics for plotting the data points. Ignored if show.data == FALSE.
...	additional arguments to be passed to the function <a href="#">contour</a> .

**Details**

contour.angmcmc is an S3 function for angmcmc objects that calls [contour](#) from graphics.

To estimate the mixture density required to construct the contour plot, first the parameter vector  $\eta$  is estimated by applying fn on the MCMC samples, yielding the (consistent) Bayes estimate  $\hat{\eta}$ . Then the mixture density  $f(x|\eta)$  at any point  $x$  is (consistently) estimated by  $f(x|\hat{\eta})$ .

**Examples**

```
# first fit a vmsin mixture model
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)
# now create a contour plot with the default first 1/3 as burn-in and thin = 1
contour(fit.vmsin.20)
```

---

densityplot1d

*Density curve for angmcmc object with univariate data*


---

**Description**

Density curve for angmcmc object with univariate data

**Usage**

```
densityplot1d(object, fn = mean, show.hist = TRUE, burnin = 1/3,
              thin = 1, ...)
```

**Arguments**

object	angular MCMC object (with univariate data).
fn	function to evaluate on MCMC samples to estimate parameters. Defaults to mean, which computes the estimated posterior mean.
show.hist	logical. Should a histogram for the data points be added to the plot?
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in $[0, 1)$ .
thin	positive integer. If $\text{thin} = n$ , only every $n$ -th realizations of the Markov chain is kept.
...	other arguments to be passed to the function <code>hist</code> from graphics. Ignored if <code>show.hist == FALSE</code> .

**Details**

To estimate the mixture density, first the parameter vector  $\eta$  is estimated by applying `fn` on the MCMC samples, yielding the (consistent) Bayes estimate  $\hat{\eta}$ . Then the mixture density  $f(x|\eta)$  at any point  $x$  is (consistently) estimated by  $f(x|\hat{\eta})$ .



**Examples**

```
# first fit a vm mixture model
# illustration only - more iterations needed for convergence
fit.vm.20 <- fit_vmmix(wind, ncomp = 2, n.iter = 20,
                     ncores = 1)
# now create density curve with the default first 1/3 as burn-in and thin = 1
densityplot1d(fit.vm.20)
```

densityplot2d

*Density surface for angmcmc objects with bivariate data***Description**

Density surface for angmcmc objects with bivariate data

**Usage**

```
densityplot2d(object, fn = mean, burnin = 1/3, thin = 1,
              log.density = FALSE, theta = 30, phi = 30, shade = 0.01,
              expand = 0.5, ...)
```

**Arguments**

object	angular MCMC object (with bivariate data).
fn	function to evaluate on MCMC samples to estimate parameters. Defaults to mean, which computes the estimated posterior mean.
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in [0, 1).
thin	positive integer. If thin = $n$ , only every $n$ -th realizations of the Markov chain is kept.
log.density	logical. Should log density be used for the plot?
theta, phi, shade, expand	arguments passed to <a href="#">persp</a> from <a href="#">graphics</a> .
...	additional arguments passed to <a href="#">persp</a> from <a href="#">graphics</a> .

**Details**

densityplot2d is a wrapper for [persp](#) from [graphics](#) applied on angmcmc objects.

To estimate the mixture density, first the parameter vector  $\eta$  is estimated by applying `fn` on the MCMC samples, yielding the (consistent) Bayes estimate  $\hat{\eta}$ . Then the mixture density  $f(x|\eta)$  at any point  $x$  is (consistently) estimated by  $f(x|\hat{\eta})$ .

**Examples**

```
# first fit a vmsin mixture model
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)
# now create density surface with the default first 1/3 as burn-in and thin = 1
densityplot2d(fit.vmsin.20)
# the viewing angles can be changed through the arguments theta and phi
# (passed to persp from graphics)
densityplot2d(fit.vmsin.20, theta = 45, phi = 45)
```

DIC

*Deviance Information Criterion (DIC) for angmcmc objects***Description**

Deviance Information Criterion (DIC) for angmcmc objects

**Usage**

```
DIC(object, form = 1, burnin = 1/3, thin = 1, ...)
```

**Arguments**

<code>object</code>	angular MCMC object.
<code>form</code>	form of DIC to use. Available choices are 1 (default) and 2. See details.
<code>burnin</code>	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in $[0, 1)$ .
<code>thin</code>	positive integer. If <code>thin = n</code> , only every $n$ -th realizations of the Markov chain is kept.
<code>...</code>	additional model specific arguments to be passed to DIC. For example, <code>int.displ</code> specifies integer displacement in <code>wnorm</code> and <code>wnorm2</code> models. See <a href="#">fit_wnormmix</a> and <a href="#">fit_wnorm2mix</a> for more details.

**Details**

Given a deviance function  $D(\theta) = -2\log(p(y|\theta))$ , and an estimate  $\theta_* = (\sum \theta_i)/N$  of the posterior mean  $E(\theta|y)$ , where  $y$  denote the data,  $\theta$  are the unknown parameters of the model,  $\theta_1, \dots, \theta_N$  are MCMC samples from the posterior distribution of  $\theta$  given  $y$  and  $p(y|\theta)$  is the likelihood function, the (form 1 of) Deviance Information Criterion (DIC) is defined as

$$DIC = 2\left(\frac{\sum_{s=1}^N D(\theta_s)}{N} - D(\theta_*)\right)$$

The second form for DIC is given by

$$DIC = D(\theta_*) - 4\hat{\text{var}} \log p(y|\theta_s)$$

where for  $i = 1, \dots, n$ ,  $\hat{v}ar \log p(y|\theta)$  denotes the estimated variance of the log likelihood based on the realizations  $\theta_1, \dots, \theta_N$ .

Like AIC and BIC, DIC is an asymptotic approximation for large samples, and is only valid when the posterior distribution is approximately normal.

### Value

Computes the DIC for a given `angmcmc` object

### Examples

```
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)
DIC(fit.vmsin.20)
```

---

<code>d_fitted</code>	<i>Density and random deviates from an <code>angmcmc</code> object</i>
-----------------------	--

---

### Description

Density and random deviates from an `angmcmc` object

### Usage

```
d_fitted(x, object, fn = mean, burnin = 1/3, thin = 1)
```

```
r_fitted(n, object, fn = mean, burnin = 1/3, thin = 1)
```

### Arguments

<code>x</code>	vector (if univariate) or a two column matrix (if bivariate, with each row a 2-D vector) of points where the densities are to be computed.
<code>object</code>	angular MCMC object. The dimension of the model must match with <code>x</code> .
<code>fn</code>	function to evaluate on MCMC samples to estimate parameters. Defaults to <code>mean</code> , which computes the estimated posterior mean.
<code>burnin</code>	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in $[0, 1)$ .
<code>thin</code>	positive integer. If <code>thin = n</code> , only every $n$ -th realizations of the Markov chain is kept.
<code>n</code>	number of observations to be generated.

**Details**

To estimate the mixture density, first the parameter vector  $\eta$  is estimated by applying `fn` on the MCMC samples (using the function `pointest`), yielding the (consistent) Bayes estimate  $\hat{\eta}$ . Then the mixture density  $f(x|\eta)$  at any point  $x$  is (consistently) estimated by  $f(x|\hat{\eta})$ .

The random deviates are generated from the estimated mixture density  $f(x|\hat{\eta})$ .

**Value**

`d_fitted` gives a vector the densities computed at the given points and `r_fitted` creates a vector (if univariate) or a matrix (if bivariate) with each row being a 2-D point, of random deviates.

**Examples**

```
set.seed(1)
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)
d_fitted(c(0,0), fit.vmsin.20)
r_fitted(10, fit.vmsin.20)
```

---

extractsamples

*Extract MCMC samples for parameters from an angmcmc object*

---

**Description**

Extract MCMC samples for parameters from an `angmcmc` object

**Usage**

```
extractsamples(object, par.name = NULL, comp.label = NULL,
               burnin = 1/3, thin = 1)
```

**Arguments**

<code>object</code>	angular MCMC object
<code>par.name</code>	vector of names of parameters for which point estimates are to be computed. If <code>NULL</code> , results for all parameters are provided.
<code>comp.label</code>	vector of component labels (positive integers, e.g., 1, 2, ...) for which point estimates are to be computed. If <code>NULL</code> , results for all components are provided.
<code>burnin</code>	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in $[0, 1)$ .
<code>thin</code>	positive integer. If <code>thin = n</code> , only every $n$ -th realizations of the Markov chain is kept.

**Details**

The default for both `par.name` and `comp.label` are the all possible choices available in `object`.

**Value**

Returns a matrix (vector) if one (both) of `par.name` and `comp.label` is of length 1, and a three dimensional array with third dimension corresponding to iterations if both `par.name` and `comp.label` have length >1.

**Examples**

```
# first fit a vmsin mixture model
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)
# extract Markov chain realizations for kappa1 from component 1
extr_kappa1_1 <- extractsamples(fit.vmsin.20, "kappa1", 1)
# for kappa1 from component from all components
extr_kappa1 <- extractsamples(fit.vmsin.20, "kappa1")
# for all parameters in component 1
extr_1 <- extractsamples(fit.vmsin.20, comp.label = 1)
```

---

fit\_stepwise\_bivariate

*Component size selection via stepwise incremented bivariate mixture models*

---

**Description**

Component size selection via stepwise incremented bivariate mixture models

**Usage**

```
fit_stepwise_bivariate(data, model, fn = mean, start_ncomp = 1,
                      max_ncomp = 10, crit = "WAIC", form = 1,
                      prev.par = FALSE, burnin = 1/3, thin = 1, ...)
```

**Arguments**

<code>data</code>	two column matrix (each row being a bivariate vector) of observations. If outside, the values are transformed into the scale $[0, 2\pi)$ .
<code>model</code>	bivariate angular mixture model to be fitted. One of "vmsin", "vmcos", or "wnorm2".
<code>fn</code>	function to evaluate on MCMC samples to estimate parameters. Defaults to mean, which computes the estimated posterior means. Ignored if <code>prev.par</code> is FALSE.
<code>start_ncomp</code>	starting component size. A single component model is fitted if <code>start_ncomp</code> is equal to one.
<code>max_ncomp</code>	maximum number of components allowed in the mixture model.

crit	criteria for model comparison, one of "AIC", "BIC", "DIC" or "WAIC". Default is "WAIC".
form	form of crit to be used. Available choices are 1 and 2. Used only if crit is "WAIC" or "DIC" and ignored otherwise.
prev.par	logical. Should the final parameters from the model with ncomp = K be used in the model with ncomp = K+1 as the starting parameters?
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in [0, 1).
thin	positive integer. If thin = n, only every n-th realizations of the Markov chain is kept.
...	additional arguments passed to <a href="#">fit_vmsinmix</a> or <a href="#">fit_vmcosmix</a> or <a href="#">fit_wnorm2mix</a> , depending on model.

### Details

The default method is "hmc". Each fit uses the number of MCMC iterations specified by the argument `n.iter` (passed to [fit\\_vmsinmix](#) or [fit\\_vmcosmix](#) or [fit\\_wnorm2mix](#), as specified through the argument `model`).

### Value

Returns a named list with the following seven elements:

`fit.all` - a list all `angmcmc` objects created at each component size;

`fit.best` - `angmcmc` object corresponding to the optimum component size;

`ncomp.best` - optimum component size (integer);

`crit` - which model comparison criterion used (one of "AIC", "BIC", "DIC" or "WAIC");

`crit.all` - all `crit` values calculated (for all component sizes);

`crit.best` - `crit` value for the optimum component size; and

`check_min` - logical; is the optimum component size less than `max_ncomp`?

### Examples

```
# illustration only - more iterations needed for convergence
fit.vmsin.step.15 <- fit_stepwise_bivariate(tim8, "vmsin", start_ncomp = 3,
                                         max_ncomp = 5, n.iter = 15,
                                         ncores = 1)

(fit.vmsin.best.15 <- bestmodel(fit.vmsin.step.15))
contour(fit.vmsin.best.15)
```

---

fit\_stepwise\_univariate

*Component size selection via stepwise incremented univariate mixture models*

---

## Description

Component size selection via stepwise incremented univariate mixture models

## Usage

```
fit_stepwise_univariate(data, model, fn = mean, start_ncomp = 1,
                        max_ncomp = 10, crit = "WAIC", form = 1,
                        prev.par = FALSE, burnin = 1/3, thin = 1, ...)
```

## Arguments

data	vector of observations. If outside, the values are transformed into the scale $[0, 2\pi)$ .
model	univariate angular mixture model to be fitted. Must be one of "vm" or "wnorm".
fn	function to evaluate on MCMC samples to estimate parameters. Defaults to mean, which computes the estimated posterior means. Ignored if prev.par is FALSE.
start_ncomp	starting component size. A single component model is fitted if start_ncomp is equal to one.
max_ncomp	maximum number of components allowed in the mixture model.
crit	criteria for model comparison, one of "AIC", "BIC", "DIC" or "WAIC". Default is "WAIC".
form	form of crit to be used. Available choices are 1 and 2. Used only if crit is "WAIC" or "DIC" and ignored otherwise.
prev.par	logical. Should the final parameters from the model with ncomp = K be used in the model with ncomp = K+1 as the starting parameters?
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in $[0, 1)$ .
thin	positive integer. If thin = n, only every n-th realizations of the Markov chain is kept.
...	additional arguments passed to <a href="#">fit_vmmix</a> or <a href="#">fit_wnormmix</a> , depending on model.

## Details

The default method is "hmc". Each fit uses the number of MCMC iterations specified by the argument n.iter (passed to [fit\\_vmmix](#) or [fit\\_wnormmix](#), as specified through the argument model).

**Value**

Returns a named list with the following seven elements:

fit.all - a list all angmcmc objects created at each component size;

fit.best - angmcmc object corresponding to the optimum component size;

ncomp.best - optimum component size (integer);

crit - which model comparison criterion used (one of "AIC", "BIC", "DIC" or "WAIC");

crit.all - all crit values calculated (for all component sizes);

crit.best - crit value for the optimum component size; and

check\_min - logical; is the optimum component size less than max\_ncomp?

**Examples**

```
# illustration only - more iterations needed for convergence
fit.vm.step.15 <- fit_stepwise_univariate(wind, "vm", start_ncomp = 1,
                                         max_ncomp = 3, n.iter = 15,
                                         ncores = 1)

(fit.vm.best.15 <- bestmodel(fit.vm.step.15))
densityplot1d(fit.vm.best.15)
```

---

fit\_vmcosmix

*Fitting bivariate von Mises cosine model mixtures using MCMC*


---

**Description**

Fitting bivariate von Mises cosine model mixtures using MCMC

**Usage**

```
fit_vmcosmix(data, ncomp, start_par = list(), method = "hmc",
             epsilon = 0.01, L = 10, epsilon.random = TRUE, L.random = FALSE,
             propscale = rep(0.01, 5), n.iter = 500, gam.loc = 0, gam.scale = 1000,
             pmix.alpha = 1/2, norm.var = 1000, autotune = FALSE, iter.tune = 10,
             ncores, show.progress = TRUE)
```

**Arguments**

data	vector of observations. If outside, the values are transformed into the scale $[0, 2\pi)$ .
ncomp	number of components in the mixture model. If comp == 1, a single component model is fitted.
start_par	list with elements pmix (ignored if comp == 1), kappa1, kappa2, mu1 and mu2, all being vectors of length same as ncomp, providing the starting values; with $j$ -th component of each vector corresponding to the $j$ -th component of the mixture distribution. If missing, moment estimators based on random components are used.



method	MCMC strategy to be used for the model parameters: "hmc" or "rwmh".
epsilon, L	tuning parameters for HMC; ignored if method = "rwmh". epsilon (step-size) is a quantity in $[0, 1)$ and L (leapfrog steps) is a positive integer.
epsilon.random	logical. Should a random value from Uniform(0, epsilon) be used for epsilon at each iteration? Ignored if method = "rwmh".
L.random	logical. Should a random value from discrete Uniform(1,..., L) be used for L at each iteration? Ignored if method = "rwmh".
propscale	tuning parameters for RWMH; a vector of size 5 representing the variances for the proposal normal densities for $\kappa_1, \kappa_2, \kappa_3, \mu_1$ and $\mu_2$ respectively. Ignored if method = "hmc".
n.iter	number of iterations for the Markov Chain.
gam.loc, gam.scale	location and scale (hyper-) parameters for the gamma prior for kappa1 and kappa2. See <a href="#">dgamma</a> . Defaults are gam.loc = 0, gam.scale = 1000 that makes the prior non-informative.
pmix.alpha	concentration parameter(s) for the Dirichlet prior for pmix. Must either be a positive real number, or a vector with positive entries and of the same size as pmix. The default is 1/2 which corresponds to the Jeffreys prior.
norm.var	variance (hyper-) parameter in the normal prior for kappa3. (Prior mean is zero). Default is 1000 that makes the prior non-informative.
autotune	logical. Should the Markov chain auto-tune the parameters (epsilon in HMC and propscale in RWMH) on the basis of acceptances in iter.tune? (The iterations used to tune the parameters are discarded.) Set to FALSE by default.
iter.tune	number of initial iterations used to tune the parameters (epsilon in HMC and propscale in RWMH). Default is 20. Ignored if autotune == FALSE.
ncores	number of CPU cores to be used for computing the likelihood, the posterior weight matrix for the Gibbs Sampler of mixture proportions and the gradient in HMC in parallel. Default is all of the available cores (obtained via <a href="#">detectCores</a> ). Note that parallelization is implemented using OpenMP. This argument is ignored and the computations are done serially if OpenMP is not available.
show.progress	logical. Should a progress bar be included?

## Details

fit\_vmcsmix generates MCMC samples of from vmcos mixture model parameters, and returns an angmcmc object as the output, which can be used as an argument for diagnostics and estimation functions.

Default method is "hmc".

If the acceptance rate drops below 5% after 100 or more HMC iterations, epsilon is automatically lowered, and the Markov chain is restarted at the current parameter values.

## Value

returns an angmcmc object.

**Examples**

```
# illustration only - more iterations needed for convergence
fit.vmcos.10 <- fit_vmcosmix(tim8, ncomp = 3, n.iter = 10,
                           ncores = 1)

fit.vmcos.10
```

fit\_vmmix

*Fitting univariate von Mises mixtures using MCMC***Description**

Fitting univariate von Mises mixtures using MCMC

**Usage**

```
fit_vmmix(data, ncomp, start_par = list(), method = "hmc", epsilon = 0.07,
          L = 10, epsilon.random = TRUE, L.random = FALSE, propscale = rep(0.01,
          2), n.iter = 500, gam.loc = 0, gam.scale = 1000, pmix.alpha = 1/2,
          autotune = FALSE, iter.tune = 10, ncores, show.progress = TRUE)
```

**Arguments**

data	vector of observations (in radians). If outside, the values are transformed into the scale $[0, 2\pi)$ .
ncomp	number of components in the mixture model. If <code>comp == 1</code> , a single component model is fitted.
start_par	list with elements <code>pmix</code> (ignored if <code>comp == 1</code> ), <code>kappa</code> and <code>mu</code> , all being vectors of length same as <code>ncomp</code> , providing the starting values; with $j$ -th component of each vector corresponding to the $j$ -th component of the mixture distribution. If missing, moment estimators based on random components are used.
method	MCMC strategy to be used for the model parameters: "hmc" or "rwmh".
epsilon, L	tuning parameters for HMC; ignored if <code>method = "rwmh"</code> . <code>epsilon</code> (step-size) is a quantity in $[0, 1)$ and <code>L</code> (leapfrog steps) is a positive integer.
epsilon.random	logical. Should a random value from <code>Uniform(0, epsilon)</code> be used for <code>epsilon</code> at each iteration? Ignored if <code>method = "rwmh"</code> .
L.random	logical. Should a random value from discrete <code>Uniform(1,..., L)</code> be used for <code>L</code> at each iteration? Ignored if <code>method = "rwmh"</code> .
propscale	tuning parameters for RWMH; a vector of size 2 representing the variances for the proposal normal densities for $\kappa$ and $\mu$ respectively. Ignored if <code>method = "hmc"</code> .
n.iter	number of iterations for the Markov Chain.
gam.loc, gam.scale	location and scale (hyper-) parameters for the gamma prior for <code>kappa</code> . See <a href="#">dgamma</a> . Defaults are <code>gam.loc = 0</code> , <code>gam.scale = 1000</code> that makes the prior non-informative.

pmix.alpha	concentration parameter(s) for the Dirichlet prior for pmix. Must either be a positive real number, or a vector with positive entries and of the same size as pmix. The default is 1/2 which corresponds to the Jeffreys prior.
autotune	logical. Should the Markov chain auto-tune the parameters (epsilon in HMC and propscale in RWMH) on the basis of acceptances in iter.tune? (The iterations used to tune the parameters are discarded.) Set to FALSE by default.
iter.tune	number of initial iterations used to tune the parameters (epsilon in HMC and propscale in RWMH). Default is 20. Ignored if autotune == FALSE.
ncores	number of CPU cores to be used for computing the likelihood, the posterior weight matrix for the Gibbs Sampler of mixture proportions and the gradient in HMC in parallel. Default is all of the available cores (obtained via <a href="#">detectCores</a> ). Note that parallelization is implemented using OpenMP. This argument is ignored and the computations are done serially if OpenMP is not available.
show.progress	logical. Should a progress bar be included?

### Details

fit\_vmmix generates MCMC samples of vm mixture model parameters, and returns an angmcmc object as the output, which can be used as an argument for diagnostics and estimation functions.

Default method is "hmc".

If the acceptance rate drops below 5% after 100 or more HMC iterations, epsilon is automatically lowered, and the Markov chain is restarted at the current parameter values.

### Value

returns an angular MCMC object.

### Examples

```
# illustration only - more iterations needed for convergence
fit.vm.20 <- fit_vmmix(wind, ncomp = 3, n.iter = 20,
                      ncores = 1)

fit.vm.20
```

---

fit\_vmsinmix

*Fitting bivariate von Mises sine model mixtures using MCMC*

---

### Description

Fitting bivariate von Mises sine model mixtures using MCMC

### Usage

```
fit_vmsinmix(data, ncomp, start_par = list(), method = "hmc",
             epsilon = 0.01, L = 10, epsilon.random = TRUE, L.random = FALSE,
             propscale = rep(0.01, 5), n.iter = 500, gam.loc = 0, gam.scale = 1000,
             pmix.alpha = 1/2, norm.var = 1000, autotune = FALSE, iter.tune = 10,
             ncores, show.progress = TRUE)
```

**Arguments**

data	vector of observations. If outside, the values are transformed into the scale $[0, 2\pi)$ .
ncomp	number of components in the mixture model. If <code>comp == 1</code> , a single component model is fitted.
start_par	list with elements <code>pmix</code> (ignored if <code>comp == 1</code> ), <code>kappa1</code> , <code>kappa2</code> , <code>mu1</code> and <code>mu2</code> , all being vectors of length same as <code>ncomp</code> , providing the starting values; with $j$ -th component of each vector corresponding to the $j$ -th component of the mixture distribution. If missing, moment estimators based on random components are used.
method	MCMC strategy to be used for the model parameters: "hmc" or "rwmh".
epsilon, L	tuning parameters for HMC; ignored if <code>method = "rwmh"</code> . <code>epsilon</code> (step-size) is a quantity in $[0, 1)$ and <code>L</code> (leapfrog steps) is a positive integer.
epsilon.random	logical. Should a random value from <code>Uniform(0, epsilon)</code> be used for <code>epsilon</code> at each iteration? Ignored if <code>method = "rwmh"</code> .
L.random	logical. Should a random value from discrete <code>Uniform(1,..., L)</code> be used for <code>L</code> at each iteration? Ignored if <code>method = "rwmh"</code> .
propscale	tuning parameters for RWMH; a vector of size 5 representing the variances for the proposal normal densities for $\kappa_1, \kappa_2, \kappa_3, \mu_1$ and $\mu_2$ respectively. Ignored if <code>method = "hmc"</code> .
n.iter	number of iterations for the Markov Chain.
gam.loc, gam.scale	location and scale (hyper-) parameters for the gamma prior for <code>kappa1</code> and <code>kappa2</code> . See <a href="#">dgamma</a> . Defaults are <code>gam.loc = 0</code> , <code>gam.scale = 1000</code> that makes the prior non-informative.
pmix.alpha	concentration parameter(s) for the Dirichlet prior for <code>pmix</code> . Must either be a positive real number, or a vector with positive entries and of the same size as <code>pmix</code> . The default is $1/2$ which corresponds to the Jeffreys prior.
norm.var	variance (hyper-) parameter in the normal prior for <code>kappa3</code> . (Prior mean is zero). Default is 1000 that makes the prior non-informative.
autotune	logical. Should the Markov chain auto-tune the parameters ( <code>epsilon</code> in HMC and <code>propscale</code> in RWMH) on the basis of acceptances in <code>iter.tune</code> ? (The iterations used to tune the parameters are discarded.) Set to <code>FALSE</code> by default.
iter.tune	number of initial iterations used to tune the parameters ( <code>epsilon</code> in HMC and <code>propscale</code> in RWMH). Default is 20. Ignored if <code>autotune == FALSE</code> .
ncores	number of CPU cores to be used for computing the likelihood, the posterior weight matrix for the Gibbs Sampler of mixture proportions and the gradient in HMC in parallel. Default is all of the available cores (obtained via <a href="#">detectCores</a> ). Note that parallelization is implemented using OpenMP. This argument is ignored and the computations are done serially if OpenMP is not available.
show.progress	logical. Should a progress bar be included?

**Details**

fit\_vmsinmix generates MCMC samples of vmsin mixture model parameters, and returns an angmcmc object as the output, which can be used as an argument for diagnostics and estimation functions.

Default method is "hmc".

If the acceptance rate drops below 5% after 100 or more HMC iterations, epsilon is automatically lowered, and the Markov chain is restarted at the current parameter values.

**Value**

returns an angmcmc object.

**Examples**

```
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)

fit.vmsin.20
```

---

fit\_wnorm2mix

*Fitting bivariate wrapped normal mixtures using MCMC*


---

**Description**

Fitting bivariate wrapped normal mixtures using MCMC

**Usage**

```
fit_wnorm2mix(data, ncomp, start_par = list(), method = "hmc",
              epsilon = 0.005, L = 10, epsilon.random = TRUE, L.random = FALSE,
              propscale = rep(0.01, 5), n.iter = 500, int.displ, gam.loc = 0,
              gam.scale = 1000, norm.var = 1000, pmix.alpha = 1/2, autotune = FALSE,
              iter.tune = 10, ncores, show.progress = TRUE)
```

**Arguments**

data	vector of observations. If outside, the values are transformed into the scale $[0, 2\pi)$ .
ncomp	number of components in the mixture model. If comp == 1, a single component model is fitted.
start_par	list with elements pmix (ignored if comp == 1), kappa1, kappa2, mu1 and mu2, all being vectors of length same as ncomp, providing the starting values; with $j$ -th component of each vector corresponding to the $j$ -th component of the mixture distribution. If missing, moment estimators based on random components are used.

method	MCMC strategy to be used for the model parameters: "hmc" or "rwmh".
epsilon, L	tuning parameters for HMC; ignored if method = "rwmh". epsilon (step-size) is a quantity in $[0, 1)$ and L (leapfrog steps) is a positive integer.
epsilon.random	logical. Should a random value from Uniform(0, epsilon) be used for epsilon at each iteration? Ignored if method = "rwmh".
L.random	logical. Should a random value from discrete Uniform(1,..., L) be used for L at each iteration? Ignored if method = "rwmh".
propscale	tuning parameters for RWMH; a vector of size 5 representing the variances for the proposal normal densities for $\kappa_1, \kappa_2, \kappa_3, \mu_1$ and $\mu_2$ respectively. Ignored if method = "hmc".
n.iter	number of iterations for the Markov Chain.
int.displ	absolute integer displacement for each coordinate. Default is 3. Allowed minimum and maximum are 1 and 5 respectively.
gam.loc, gam.scale	location and scale (hyper-) parameters for the gamma prior for kappa1 and kappa2. See <a href="#">dgamma</a> . Defaults are gam.loc = 0, gam.scale = 1000 that makes the prior non-informative.
norm.var	variance (hyper-) parameter in the normal prior for kappa3. (Prior mean is zero). Default is 1000 that makes the prior non-informative.
pmix.alpha	concentration parameter(s) for the Dirichlet prior for pmix. Must either be a positive real number, or a vector with positive entries and of the same size as pmix. The default is 1/2 which corresponds to the Jeffreys prior.
autotune	logical. Should the Markov chain auto-tune the parameters (epsilon in HMC and propscale in RWMH) on the basis of acceptances in iter.tune? (The iterations used to tune the parameters are discarded.) Set to FALSE by default.
iter.tune	number of initial iterations used to tune the parameters (epsilon in HMC and propscale in RWMH). Default is 20. Ignored if autotune == FALSE.
ncores	number of CPU cores to be used for computing the likelihood, the posterior weight matrix for the Gibbs Sampler of mixture proportions and the gradient in HMC in parallel. Default is all of the available cores (obtained via <a href="#">detectCores</a> ). Note that parallelization is implemented using OpenMP. This argument is ignored and the computations are done serially if OpenMP is not available.
show.progress	logical. Should a progress bar be included?

## Details

fit\_wnorm2mix generates MCMC samples of wnorm2 mixture model parameters, and returns an angmcmc object as the output, which can be used as an argument for diagnostics and estimation functions.

Default method is "hmc".

If the acceptance rate drops below 5% after 100 or more HMC iterations, epsilon is automatically lowered, and the Markov chain is restarted at the current parameter values.

**Value**

returns an `angmcmc` object.

**Examples**

```
# illustration only - more iterations needed for convergence
fit.wnorm2.15 <- fit_wnorm2mix(tim8, ncomp = 3, n.iter = 15,
                             ncores = 1)

fit.wnorm2.15
```

---

 fit\_wnormmix

*Fitting univariate wrapped normal mixtures using MCMC*


---

**Description**

Fitting univariate wrapped normal mixtures using MCMC

**Usage**

```
fit_wnormmix(data, ncomp, start_par, method = "hmc", epsilon = 0.07,
             L = 10, epsilon.random = TRUE, L.random = FALSE, propscale = rep(0.01,
             2), n.iter = 10000, int.displ, gam.loc = 0, gam.scale = 1000,
             pmix.alpha = 1/2, autotune = FALSE, iter.tune = 10, ncores,
             show.progress = TRUE)
```

**Arguments**

<code>data</code>	vector of observations (in radians). If outside, the values are transformed into the scale $[0, 2\pi)$ .
<code>ncomp</code>	number of components in the mixture model. If <code>comp == 1</code> , a single component model is fitted.
<code>start_par</code>	list with elements <code>pmix</code> (ignored if <code>comp == 1</code> ), <code>kappa</code> and <code>mu</code> , all being vectors of length same as <code>ncomp</code> , providing the starting values; with $j$ -th component of each vector corresponding to the $j$ -th component of the mixture distribution. If missing, moment estimators based on random components are used.
<code>method</code>	MCMC strategy to be used for the model parameters: "hmc" or "rwmh".
<code>epsilon, L</code>	tuning parameters for HMC; ignored if <code>method = "rwmh"</code> . <code>epsilon</code> (step-size) is a quantity in $[0, 1)$ and <code>L</code> (leapfrog steps) is a positive integer.
<code>epsilon.random</code>	logical. Should a random value from <code>Uniform(0, epsilon)</code> be used for <code>epsilon</code> at each iteration? Ignored if <code>method = "rwmh"</code> .
<code>L.random</code>	logical. Should a random value from discrete <code>Uniform(1, ..., L)</code> be used for <code>L</code> at each iteration? Ignored if <code>method = "rwmh"</code> .
<code>propscale</code>	tuning parameters for RWMH; a vector of size 2 representing the variances for the proposal normal densities for $\kappa$ and $\mu$ respectively. Ignored if <code>method = "hmc"</code> .

n.iter	number of iterations for the Markov Chain.
int.displ	integer displacement. The allowed values are 1, 2, 3, 4 and 5. Default is 3.
gam.loc, gam.scale	location and scale (hyper-) parameters for the gamma prior for kappa. See <a href="#">dgamma</a> . Defaults are gam.loc = 0, gam.scale = 1000 that makes the prior non-informative.
pmix.alpha	concentration parameter(s) for the Dirichlet prior for pmix. Must either be a positive real number, or a vector with positive entries and of the same size as pmix. The default is 1/2 which corresponds to the Jeffreys prior.
autotune	logical. Should the Markov chain auto-tune the parameters (epsilon in HMC and propscale in RWMH) on the basis of acceptances in iter.tune? (The iterations used to tune the parameters are discarded.) Set to FALSE by default.
iter.tune	number of initial iterations used to tune the parameters (epsilon in HMC and propscale in RWMH). Default is 20. Ignored if autotune == FALSE.
ncores	number of CPU cores to be used for computing the likelihood, the posterior weight matrix for the Gibbs Sampler of mixture proportions and the gradient in HMC in parallel. Default is all of the available cores (obtained via <a href="#">detectCores</a> ). Note that parallelization is implemented using OpenMP. This argument is ignored and the computations are done serially if OpenMP is not available.
show.progress	logical. Should a progress bar be included?

## Details

fit\_wnormmix generates MCMC samples of wnorm mixture model parameters, and returns an angmcmc object as the output, which can be used as an argument for diagnostics and estimation functions.

Default method is "hmc".

If the acceptance rate drops below 5% after 100 or more HMC iterations, epsilon is automatically lowered, and the Markov chain is restarted at the current parameter values.

## Value

returns an angular MCMC object.

## Examples

```
# illustration only - more iterations needed for convergence
fit.wnorm.20 <- fit_wnormmix(wind, ncomp = 3, n.iter = 20,
                           ncores = 1)

fit.wnorm.20
```



---

`fix_label`*Fix label switching in angmcmc objects*

---

## Description

Fix label switching in angmcmc objects

## Usage

```
fix_label(object, burnin = 1/3, thin = 1, method = 1)
```

## Arguments

<code>object</code>	angular MCMC object.
<code>burnin</code>	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in $[0, 1)$ .
<code>thin</code>	positive integer. If <code>thin = n</code> , only every $n$ -th realizations of the Markov chain is kept.
<code>method</code>	method to use for fixing label switching. Available choices are 1 for <a href="#">dataBased</a> and 2 for <a href="#">ecr.iterative.1</a> .

## Details

`fix_label` is a wrapper for [dataBased](#) or [ecr.iterative.1](#) (depending on `method`) from the `label.switching` package for angmcmc objects.

## Value

Returns another angmcmc object, with the parameter values (after burn-in and thin) re-ordered according to the resulting permutation from [dataBased](#) or [ecr.iterative.1](#).

## Examples

```
# first fit a vmsin mixture model
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                          ncores = 1)

# now apply fix_label
fit.vmsin.20.fix <- fix_label(fit.vmsin.20)
```





---

lpdtrace	<i>Trace plot of log posterior density or log likelihood from an angmcmc object</i>
----------	---

---

**Description**

Trace plot of log posterior density or log likelihood from an angmcmc object

**Usage**

```
lpdtrace(object, use.llik = FALSE, burnin = 1/3, thin = 1)
```

**Arguments**

object	angular MCMC object.
use.llik	logical. Should log likelihood be plotted instead of log posterior? Set to FALSE by default.
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in [0, 1).
thin	positive integer. If thin = $n$ , only every $n$ -th realizations of the Markov chain is kept.

**Examples**

```
# first fit a vmsin mixture model
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)
# log posterior density trace
lpdtrace(fit.vmsin.20)
# log likelihood trace
lpdtrace(fit.vmsin.20, use.llik = TRUE)
```

---

paramtrace	<i>Trace plot for parameters from an angmcmc object</i>
------------	---

---

**Description**

Trace plot for parameters from an angmcmc object

**Usage**

```
paramtrace(object, par, comp.label, burnin = 1/3, thin = 1)
```

**Arguments**

object	angular MCMC object.
par	parameter for which trace plot is to be created.
comp.label	vector of component labels (positive integers, e.g., 1, 2, ...) for which point estimates are to be computed. If NULL, results for all components are provided.
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in [0, 1).
thin	positive integer. If thin = $n$ , only every $n$ -th realizations of the Markov chain is kept.

**Value**

Returns a single plot if a single par and a single comp.label is supplied. Otherwise, a series of plots is produced.

**Examples**

```
# first fit a vmsin mixture model
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)
# trace plot for kappa1 in component 1
paramtrace(fit.vmsin.20, "kappa1", 1)
# for kappa1 in all components
paramtrace(fit.vmsin.20, "kappa1")
# for all parameters in component 1
paramtrace(fit.vmsin.20, comp.label = 1)
```

---

pointest

*Point estimates for parameters from an angmcmc object*


---

**Description**

Point estimates for parameters from an angmcmc object

**Usage**

```
pointest(object, fn = mean, par.name = NULL, comp.label = NULL,
         burnin = 1/3, thin = 1)
```

**Arguments**

object	angular MCMC object.
fn	function to evaluate on MCMC samples to estimate parameters. Defaults to mean, which computes the estimated posterior mean.
par.name	vector of names of parameters for which point estimates are to be computed. If NULL, results for all parameters are provided.
comp.label	vector of component labels (positive integers, e.g., 1, 2, ...) for which point estimates are to be computed. If NULL, results for all components are provided.
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in [0, 1).
thin	positive integer. If thin = n, only every n-th realizations of the Markov chain is kept.

**Value**

Returns a matrix of point estimates, or vector of point estimates if `length(par.name)==1` or `length(comp.label)==1`.

**Examples**

```
# first fit a vmsin mixture model
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)
# estimate parameters by sample mean
(est_mean <- pointest(fit.vmsin.20))
# estimate parameters by sample median
(est_median <- pointest(fit.vmsin.20, fn = median))
```

---

quantile.angmcmc

*Quantile estimates for parameters from an angmcmc object*

---

**Description**

Quantile estimates for parameters from an angmcmc object

**Usage**

```
## S3 method for class 'angmcmc'
quantile(x, par.name = NULL, comp.label = NULL,
         burnin = 1/3, thin = 1, ...)
```

**Arguments**

x	angmcmc object
par.name	vector of names of parameters for which point estimates are to be computed. If NULL, results for all parameters are provided.
comp.label	vector of component labels (positive integers, e.g., 1, 2, ...) for which point estimates are to be computed. If NULL, results for all components are provided.
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in [0, 1).
thin	positive integer. If thin = n, only every n-th realizations of the Markov chain is kept.
...	further arguments to pass to quantile. In particular, probs = seq(0, 1, 0.25) is the default vector of quantiles computed for each parameter.

**Value**

Returns a three dimensional array of quantiles, or a matrix (vector) of quantiles if one (or two) among par.name, comp.label, probs has length 1.

**Examples**

```
# first fit a vmsin mixture model
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)

# 0.025th quantiles
(quant_025 <- quantile(fit.vmsin.20, prob = 0.025))
# 0.975th quantiles
(quant_975 <- quantile(fit.vmsin.20, prob = 0.975))
# default quantiles
(quant_def <- quantile(fit.vmsin.20))
```

---

rvm

*The univariate von Mises distribution*


---

**Description**

The univariate von Mises distribution

**Usage**

```
rvm(n, kappa = 1, mu = 0)
```

```
dvm(x, kappa = 1, mu = 0)
```

**Arguments**

n	number of observations. Ignored if at least one of the other parameters have length $k > 1$ , in which case, all the parameters are recycled to length $k$ to produce $k$ random variates.
kappa	vector of concentration (inverse-variance) parameters; $\text{kappa} > 0$ .
mu	vector of means.
x	vector of angles (in radians) where the densities are to be evaluated.

**Details**

If  $\mu$  and  $\text{kappa}$  are not specified they assume the default values of 0 and 1 respectively.

The univariate von Mises distribution has density

$$f(x) = 1/(2\pi I_0(\kappa)) \exp(\kappa \cos(x - \mu))$$

where  $I_0(\kappa)$  denotes the modified Bessel function of the first kind with order 0 evaluated at the point  $\kappa$ .

**Value**

dvm gives the density and rvm generates random deviates.

**Examples**

```
kappa <- 1:3
mu <- 0:2
x <- 1:10
n <- 10

# when x and both parameters are scalars, dvm returns a single density
dvm(x[1], kappa[1], mu[1])

# when x is a vector but both the parameters are scalars, dvm returns a vector of
# densities calculated at each entry of x with the same parameters
dvm(x, kappa[1], mu[1])

# if x is scalar and at least one of the two parameters is a vector, both parameters are
# recycled to the same length, and dvm returns a vector of with ith element being the
# density evaluated at x with parameter values kappa[i] and mu[i]
dvm(x[1], kappa, mu)

# if x and at least one of the two parameters is a vector, x and the two parameters are
# recycled to the same length, and dvm returns a vector of with ith element being the
# density at ith element of the (recycled) x with parameter values kappa[i] and mu[i]
dvm(x, kappa, mu)

# when parameters are all scalars, number of observations generated by rvm is n
rvm(n, kappa[1], mu[1])
```



```
# when at least one of the two parameters is a vector, both are recycled to the same length,
# n is ignored, and the number of observations generated by rvm is the same as the length of
# the recycled vectors
rvm(n, kappa, mu)
```

---

rvmcos

*The bivariate von Mises cosine model*


---

### Description

The bivariate von Mises cosine model

### Usage

```
rvmcos(n, kappa1 = 1, kappa2 = 1, kappa3 = 0, mu1 = 0, mu2 = 0)
```

```
dvmcos(x, kappa1 = 1, kappa2 = 1, kappa3 = 0, mu1 = 0, mu2 = 0)
```

### Arguments

n	number of observations. Ignored if at least one of the other parameters have length $k > 1$ , in which case, all the parameters are recycled to length $k$ to produce $k$ random variates.
kappa1, kappa2, kappa3	vectors of concentration parameters; $\text{kappa1}, \text{kappa2} > 0$ .
mu1, mu2	vectors of mean parameters.
x	bivariate vector or a two-column matrix with each row being a bivariate vector of angles (in radians) where the densities are to be evaluated.

### Details

The bivariate von Mises cosine model density at the point  $x = (x_1, x_2)$  is given by

$$f(x) = C_c(\kappa_1, \kappa_2, \kappa_3) \exp(\kappa_1 \cos(T_1) + \kappa_2 \cos(T_2) + \kappa_3 \cos(T_1 - T_2))$$

where

$$T_1 = x_1 - \mu_1; T_2 = x_2 - \mu_2$$

and  $C_c(\kappa_1, \kappa_2, \kappa_3)$  denotes the normalizing constant for the cosine model.

### Value

dvmcos gives the density and rvmcos generates random deviates.

**Examples**

```

kappa1 <- c(1, 2, 3)
kappa2 <- c(1, 6, 5)
kappa3 <- c(0, 1, 2)
mu1 <- c(1, 2, 5)
mu2 <- c(0, 1, 3)
x <- diag(2, 2)
n <- 10

# when x is a bivariate vector and parameters are all scalars,
# dvmcos returns single density
dvmcos(x[1, ], kappa1[1], kappa2[1], kappa3[1], mu1[1], mu2[1])

# when x is a two column matrix and parameters are all scalars,
# dmvsin returns a vector of densities calculated at the rows of
# x with the same parameters
dvmcos(x, kappa1[1], kappa2[1], kappa3[1], mu1[1], mu2[1])

# if x is a bivariate vector and at least one of the parameters is
# a vector, all parameters are recycled to the same length, and
# dvmcos returns a vector with ith element being the density
# evaluated at x with parameter values kappa1[i], kappa2[i],
# kappa3[i], mu1[i] and mu2[i]
dvmcos(x[1, ], kappa1, kappa2, kappa3, mu1, mu2)

# if x is a two column matrix and at least one of the parameters is
# a vector, rows of x and the parameters are recycled to the same
# length, and dvmcos returns a vector with ith element being the
# density evaluated at ith row of x with parameter values kappa1[i],
# kappa2[i], # kappa3[i], mu1[i] and mu2[i]
dvmcos(x[1, ], kappa1, kappa2, kappa3, mu1, mu2)

# when parameters are all scalars, number of observations generated
# by rvmcos is n
rvmcos(n, kappa1[1], kappa2[1], kappa3[1], mu1[1], mu2[1])

# when at least one of the parameters is a vector, all parameters are
# recycled to the same length, n is ignored, and the number of
# observations generated by rvmcos is the same as the length of the
# recycled vectors
rvmcos(n, kappa1, kappa2, kappa3, mu1, mu2)

```

---

rvmcosmix

*The bivariate von Mises cosine model mixtures*


---

**Description**

The bivariate von Mises cosine model mixtures

**Usage**

```
rvmcosmix(n, kappa1, kappa2, kappa3, mu1, mu2, pmix)
```

```
dvmcosmix(x, kappa1, kappa2, kappa3, mu1, mu2, pmix)
```

**Arguments**

`n` number of observations.  
`kappa1, kappa2, kappa3` vectors of concentration parameters;  $\text{kappa1}, \text{kappa2} > 0$  for each component.  
`mu1, mu2` vectors of mean parameters.  
`pmix` vector of mixture proportions.  
`x` matrix of angles (in radians) where the density is to be evaluated, with each row being a single bivariate vector of angles.

**Details**

All the argument vectors `pmix`, `kappa1`, `kappa2`, `kappa3`, `mu1` and `mu2` must be of the same length (= component size of the mixture model), with  $j$ -th element corresponding to the  $j$ -th component of the mixture distribution.

The bivariate von Mises cosine model mixture distribution with component size  $K = \text{length}(\text{pmix})$  has density

$$g(x) = \sum p[j] * f(x; \kappa_1[j], \kappa_2[j], \kappa_3[j], \mu_1[j], \mu_2[j])$$

where the sum extends over  $j$ ;  $p[j]$ ;  $\kappa_1[j]$ ,  $\kappa_2[j]$ ,  $\kappa_3[j]$ ; and  $\mu_1[j]$ ,  $\mu_2[j]$  respectively denote the mixing proportion, the three concentration parameters and the two mean parameter for the  $j$ -th cluster,  $j = 1, \dots, K$ , and  $f(\cdot; \kappa_1, \kappa_2, \kappa_3, \mu_1, \mu_2)$  denotes the density function of the von Mises cosine model with concentration parameters  $\kappa_1, \kappa_2, \kappa_3$  and mean parameters  $\mu_1, \mu_2$ .

**Value**

`dvmcosmix` computes the density and `rvmcosmix` generates random deviates from the mixture density.

**Examples**

```
kappa1 <- c(1, 2, 3)
kappa2 <- c(1, 6, 5)
kappa3 <- c(0, 1, 2)
mu1 <- c(1, 2, 5)
mu2 <- c(0, 1, 3)
pmix <- c(0.3, 0.4, 0.3)
x <- diag(2, 2)
n <- 10

# mixture densities calculated at the rows of x
dvmcosmix(x, kappa1, kappa2, kappa3, mu1, mu2, pmix)

# number of observations generated from the mixture distribution is n
```

```
rvmcosmix(n, kappa1, kappa2, kappa3, mu1, mu2, pmix)
```

---

 rvmmix

*The univariate von Mises mixtures*


---

### Description

The univariate von Mises mixtures

### Usage

```
rvmmix(n, kappa, mu, pmix)
```

```
dvmmix(x, kappa, mu, pmix)
```

### Arguments

n	number of observations. Ignored if at least one of the other parameters have length $k > 1$ , in which case, all the parameters are recycled to length $k$ to produce $k$ random variates.
kappa	vector of component concentration (inverse-variance) parameters, $\text{kappa} > 0$ .
mu	vector of component means.
pmix	vector of mixing proportions.
x	vector of angles (in radians) where the densities are to be evaluated.

### Details

pmix, mu and kappa must be of the same length, with  $j$ -th element corresponding to the  $j$ -th component of the mixture distribution.

The univariate von Mises mixture distribution with component size  $K = \text{length}(\text{pmix})$  has density

$$g(x) = p[1] * f(x; \kappa[1], \mu[1]) + \dots + p[K] * f(x; \kappa[K], \mu[K])$$

where  $p[j]$ ,  $\kappa[j]$ ,  $\mu[j]$  respectively denote the mixing proportion, concentration parameter and the mean parameter for the  $j$ -th component and  $f(\cdot; \kappa, \mu)$  denotes the density function of the (univariate) von Mises distribution with mean parameter  $\mu$  and concentration parameter  $\kappa$ .

### Value

dvmmix computes the density and rvmmix generates random deviates from the mixture density.

**Examples**

```

kappa <- 1:3
mu <- 0:2
pmix <- c(0.3, 0.3, 0.4)
x <- 1:10
n <- 10

# mixture densities calculated at each point in x
dvmmix(x, kappa, mu, pmix)

# number of observations generated from the mixture distribution is n
rvmmix(n, kappa, mu, pmix)

```

---

rvmsin

*The bivariate von Mises sine model*


---

**Description**

The bivariate von Mises sine model

**Usage**

```
rvmsin(n, kappa1 = 1, kappa2 = 1, kappa3 = 0, mu1 = 0, mu2 = 0)
```

```
dvmsin(x, kappa1 = 1, kappa2 = 1, kappa3 = 0, mu1 = 0, mu2 = 0)
```

**Arguments**

**n** number of observations. Ignored if at least one of the other parameters have length  $k > 1$ , in which case, all the parameters are recycled to length  $k$  to produce  $k$  random variates.

**kappa1, kappa2, kappa3** vectors of concentration parameters;  $\text{kappa1}, \text{kappa2} > 0$ .

**mu1, mu2** vectors of mean parameters.

**x** bivariate vector or a two-column matrix with each row being a bivariate vector of angles (in radians) where the densities are to be evaluated.

**Details**

The bivariate von Mises sine model density at the point  $x = (x_1, x_2)$  is given by

$$f(x) = C_s(\kappa_1, \kappa_2, \kappa_3) \exp(\kappa_1 \cos(T_1) + \kappa_2 \cos(T_2) + \kappa_3 \sin(T_1) \sin(T_2))$$

where

$$T_1 = x_1 - \mu_1; T_2 = x_2 - \mu_2$$

and  $C_s(\kappa_1, \kappa_2, \kappa_3)$  denotes the normalizing constant for the sine model.

**Value**

dvmsin gives the density and rvmsin generates random deviates.

**Examples**

```

kappa1 <- c(1, 2, 3)
kappa2 <- c(1, 6, 5)
kappa3 <- c(0, 1, 2)
mu1 <- c(1, 2, 5)
mu2 <- c(0, 1, 3)
x <- diag(2, 2)
n <- 10

# when x is a bivariate vector and parameters are all scalars,
# dvmsin returns single density
dvmsin(x[1, ], kappa1[1], kappa2[1], kappa3[1], mu1[1], mu2[1])

# when x is a two column matrix and parameters are all scalars,
# dmvsin returns a vector of densities calculated at the rows of
# x with the same parameters
dvmsin(x, kappa1[1], kappa2[1], kappa3[1], mu1[1], mu2[1])

# if x is a bivariate vector and at least one of the parameters is
# a vector, all parameters are recycled to the same length, and
# dvmsin returns a vector of with ith element being the density
# evaluated at x with parameter values kappa1[i], kappa2[i],
# kappa3[i], mu1[i] and mu2[i]
dvmsin(x[1, ], kappa1, kappa2, kappa3, mu1, mu2)

# if x is a two column matrix and at least one of the parameters is
# a vector, rows of x and the parameters are recycled to the same
# length, and dvmsin returns a vector of with ith element being the
# density evaluated at ith row of x with parameter values kappa1[i],
# kappa2[i], # kappa3[i], mu1[i] and mu2[i]
dvmsin(x[1, ], kappa1, kappa2, kappa3, mu1, mu2)

# when parameters are all scalars, number of observations generated
# by rvmsin is n
rvmsin(n, kappa1[1], kappa2[1], kappa3[1], mu1[1], mu2[1])

# when at least one of the parameters is a vector, all parameters are
# recycled to the same length, n is ignored, and the number of
# observations generated by rvmsin is the same as the length of the
# recycled vectors
rvmsin(n, kappa1, kappa2, kappa3, mu1, mu2)

```

**Description**

The bivariate von Mises sine model mixtures

**Usage**

```
rvmsinmix(n, kappa1, kappa2, kappa3, mu1, mu2, pmix)
```

```
dvmsinmix(x, kappa1, kappa2, kappa3, mu1, mu2, pmix)
```

**Arguments**

`n` number of observations.  
`kappa1, kappa2, kappa3` vectors of concentration parameters; `kappa1, kappa2 > 0` for each component.  
`mu1, mu2` vectors of mean parameters.  
`pmix` vector of mixture proportions.  
`x` matrix of angles (in radians) where the density is to be evaluated, with each row being a single bivariate vector of angles.

**Details**

All the argument vectors `pmix, kappa1, kappa2, kappa3, mu1` and `mu2` must be of the same length (= component size of the mixture model), with  $j$ -th element corresponding to the  $j$ -th component of the mixture distribution.

The bivariate von Mises sine model mixture distribution with component size  $K = \text{length}(\text{p.mix})$  has density

$$g(x) = \sum p[j] * f(x; \kappa_1[j], \kappa_2[j], \kappa_3[j], \mu_1[j], \mu_2[j])$$

where the sum extends over  $j$ ;  $p[j]; \kappa_1[j], \kappa_2[j], \kappa_3[j]$ ; and  $\mu_1[j], \mu_2[j]$  respectively denote the mixing proportion, the three concentration parameters and the two mean parameter for the  $j$ -th component,  $j = 1, \dots, K$ , and  $f(.; \kappa_1, \kappa_2, \kappa_3, \mu_1, \mu_2)$  denotes the density function of the von Mises sine model with concentration parameters  $\kappa_1, \kappa_2, \kappa_3$  and mean parameters  $\mu_1, \mu_2$ .

**Value**

`dvmsinmix` computes the density (vector if `x` is a two column matrix with more than one row) and `rvmsinmix` generates random deviates from the mixture density.

**Examples**

```
kappa1 <- c(1, 2, 3)
kappa2 <- c(1, 6, 5)
kappa3 <- c(0, 1, 2)
mu1 <- c(1, 2, 5)
mu2 <- c(0, 1, 3)
pmix <- c(0.3, 0.4, 0.3)
x <- diag(2, 2)
n <- 10
```

```
# mixture densities calculated at the rows of x
dvmsinmix(x, kappa1, kappa2, kappa3, mu1, mu2, pmix)

# number of observations generated from the mixture distribution is n
rvmsinmix(n, kappa1, kappa2, kappa3, mu1, mu2, pmix)
```

---

rwnorm

*The univariate Wrapped Normal distribution*


---

### Description

The univariate Wrapped Normal distribution

### Usage

```
rwnorm(n = 1, kappa = 1, mu = 0)

dwnorm(x, kappa = 1, mu = 0, int.displ)
```

### Arguments

n	number of observations. Ignored if at least one of the other parameters have length $k > 1$ , in which case, all the parameters are recycled to length $k$ to produce $k$ random variates.
kappa	vector of concentration (inverse-variance) parameters; $kappa > 0$ .
mu	vector of means.
x	vector of angles (in radians) where the densities are to be evaluated.
int.displ	integer displacement. The allowed values are 1, 2, 3, 4 and 5. Default is 3.

### Details

If mu and kappa are not specified they assume the default values of 0 and 1 respectively.

The univariate wrapped normal distribution has density

$$f(x) = \sqrt{\kappa/(2\pi)} \sum \exp(-\kappa/2(x - \mu(2\pi\omega))^2)$$

where the sum extends over all integers  $\omega$ , and is approximated by a sum over  $\omega$  in  $\{-M, -M + 1, \dots, M - 1, M\}$  if  $\text{int.displ} = M$ .

### Value

dwnorm gives the density and rwnorm generates random deviates.



**Examples**

```

kappa <- 1:3
mu <- 0:2
x <- 1:10
n <- 10

# when x and both parameters are scalars, dwnorm returns a single density
dwnorm(x[1], kappa[1], mu[1])

# when x is a vector but both the parameters are scalars, dmv returns a vector of
# densities calculated at each entry of x with the same parameters
dwnorm(x, kappa[1], mu[1])

# if x is scalar and at least one of the two parameters is a vector, both parameters are
# recycled to the same length, and dwnorm returns a vector of with ith element being the
# density evaluated at x with parameter values kappa[i] and mu[i]
dwnorm(x[1], kappa, mu)

# if x and at least one of the two parameters is a vector, x and the two parameters are
# recycled to the same length, and dwnorm returns a vector of with ith element being the
# density at ith element of the (recycled) x with parameter values kappa[i] and mu[i]
dwnorm(x, kappa, mu)

# when parameters are all scalars, number of observations generated by rwnorm is n
rwnorm(n, kappa[1], mu[1])

# when at least one of the two parameters is a vector, both are recycled to the same length,
# n is ignored, and the number of observations generated by rwnorm is the same as the length
# of the recycled vectors
rwnorm(n, kappa, mu)

```

---

rwnorm2

*The bivariate Wrapped Normal distribution*


---

**Description**

The bivariate Wrapped Normal distribution

**Usage**

```
rwnorm2(n, kappa1 = 1, kappa2 = 1, kappa3 = 0, mu1 = 0, mu2 = 0)
```

```
dwnorm2(x, kappa1 = 1, kappa2 = 1, kappa3 = 0, mu1 = 0, mu2 = 0,
int.displ)
```

**Arguments**

n	number of observations. Ignored if at least one of the other parameters have length $k > 1$ , in which case, all the parameters are recycled to length $k$ to produce $k$ random variates.
kappa1, kappa2, kappa3	vectors of concentration parameters; $\text{kappa1}, \text{kappa2} > 0$ , and $\text{kappa3}^2 < \text{kappa1} * \text{kappa2}$ .
mu1, mu2	vectors of mean parameters.
x	bivariate vector or a two-column matrix with each row being a bivariate vector of angles (in radians) where the densities are to be evaluated.
int.displ	integer displacement. The allowed values are 1, 2, 3, 4 and 5. Default is 3.

**Details**

The bivariate wrapped normal density at the point  $x = (x_1, x_2)$  is given by,

$$f(x) = \sqrt{((\kappa_1 \kappa_2 - (\kappa_3)^2)) / (2\pi)} \sum \exp(-1/2 * (\kappa_1(T_1)^2 + \kappa_2(T_2)^2 + 2\kappa_3(T_1)(T_2)))$$

where

$$T_1 = T_1(x, \mu, \omega) = (x_1 - \mu_1(2\pi\omega_1))$$

$$T_2 = T_2(x, \mu, \omega) = (x_2 - \mu_2(2\pi\omega_2))$$

the sum extends over all pairs of integers  $\omega = (\omega_1, \omega_2)$ , and is approximated by a sum over  $(\omega_1, \omega_2)$  in  $\{-M, -M + 1, \dots, M - 1, M\}^2$  if `int.displ = M`.

Note that above density is essentially the "wrapped" version of a bivariate normal density with mean

$$\mu = (\mu_1, \mu_2)$$

and dispersion matrix  $\Sigma = \Delta^{-1}$ , where

$$\Delta = \begin{matrix} \kappa_1 & \kappa_3 \\ \kappa_3 & \kappa_2 \end{matrix}$$

**Value**

`dwnorm2` gives the density and `rwnorm2` generates random deviates.

**Examples**

```
kappa1 <- c(1, 2, 3)
kappa2 <- c(1, 6, 5)
kappa3 <- c(0, 1, 2)
mu1 <- c(1, 2, 5)
mu2 <- c(0, 1, 3)
x <- diag(2, 2)
n <- 10
```

# when x is a bivariate vector and parameters are all scalars,

```

# dwnorm2 returns single density
dwnorm2(x[1, ], kappa1[1], kappa2[1], kappa3[1], mu1[1], mu2[1])

# when x is a two column matrix and parameters are all scalars,
# dwnorm2 returns a vector of densities calculated at the rows of
# x with the same parameters
dwnorm2(x, kappa1[1], kappa2[1], kappa3[1], mu1[1], mu2[1])

# if x is a bivariate vector and at least one of the parameters is
# a vector, all parameters are recycled to the same length, and
# dwnorm2 returns a vector of with ith element being the density
# evaluated at x with parameter values kappa1[i], kappa2[i],
# kappa3[i], mu1[i] and mu2[i]
dwnorm2(x[1, ], kappa1, kappa2, kappa3, mu1, mu2)

# if x is a two column matrix and at least one of the parameters is
# a vector, rows of x and the parameters are recycled to the same
# length, and dwnorm2 returns a vector of with ith element being the
# density evaluated at ith row of x with parameter values kappa1[i],
# kappa2[i], # kappa3[i], mu1[i] and mu2[i]
dwnorm2(x[1, ], kappa1, kappa2, kappa3, mu1, mu2)

# when parameters are all scalars, number of observations generated
# by rwnorm2 is n
rwnorm2(n, kappa1[1], kappa2[1], kappa3[1], mu1[1], mu2[1])

# when at least one of the parameters is a vector, all parameters are
# recycled to the same length, n is ignored, and the number of
# observations generated by rwnorm2 is the same as the length of the
# recycled vectors
rwnorm2(n, kappa1, kappa2, kappa3, mu1, mu2)

```

---

rwnorm2mix

*The bivariate Wrapped Normal mixtures*


---

## Description

The bivariate Wrapped Normal mixtures

## Usage

```
rwnorm2mix(n, kappa1, kappa2, kappa3, mu1, mu2, pmix)
```

```
dwnorm2mix(x, kappa1, kappa2, kappa3, mu1, mu2, pmix, int.displ)
```

## Arguments

n                    number of observations.

kappa1, kappa2, kappa3	vectors of concentration parameters; kappa1, kappa2 > 0, kappa3^2 < kappa1*kappa2 for each component.
mu1, mu2	vectors of mean parameters.
pmix	vector of mixture proportions.
x	matrix of angles (in radians) where the density is to be evaluated, with each row being a single bivariate vector of angles.
int.displ	integer displacement. The allowed values are 1, 2, 3, 4 and 5. Default is 3.

### Details

All the argument vectors pmix, kappa1, kappa2, kappa3, mu1 and mu2 must be of the same length, with  $j$ -th element corresponding to the  $j$ -th component of the mixture distribution.

The bivariate wrapped normal mixture distribution with component size  $K = \text{length}(\text{pmix})$  has density

$$g(x) = \sum p[j] * f(x; \kappa_1[j], \kappa_2[j], \kappa_3[j], \mu_1[j], \mu_2[j])$$

where the sum extends over  $j$ ;  $p[j]$ ;  $\kappa_1[j]$ ,  $\kappa_2[j]$ ,  $\kappa_3[j]$ ; and  $\mu_1[j]$ ,  $\mu_2[j]$  respectively denote the mixing proportion, the three concentration parameters and the two mean parameter for the  $j$ -th component,  $j = 1, \dots, K$ , and  $f(\cdot; \kappa_1, \kappa_2, \kappa_3, \mu_1, \mu_2)$  denotes the density function of the wrapped normal distribution with concentration parameters  $\kappa_1, \kappa_2, \kappa_3$  and mean parameters  $\mu_1, \mu_2$ .

### Value

dwnorm2mix computes the density and rwnorm2mix generates random deviates from the mixture density.

### Examples

```
kappa1 <- c(1, 2, 3)
kappa2 <- c(1, 6, 5)
kappa3 <- c(0, 1, 2)
mu1 <- c(1, 2, 5)
mu2 <- c(0, 1, 3)
pmix <- c(0.3, 0.4, 0.3)
x <- diag(2, 2)
n <- 10

# mixture densities calculated at the rows of x
dwnorm2mix(x, kappa1, kappa2, kappa3, mu1, mu2, pmix)

# number of observations generated from the mixture distribution is n
rwnorm2mix(n, kappa1, kappa2, kappa3, mu1, mu2, pmix)
```

---

rwnormmix

*The univariate Wrapped Normal mixtures*


---

## Description

The univariate Wrapped Normal mixtures

## Usage

```
rwnormmix(n = 1, kappa, mu, pmix)
```

```
dwnormmix(x, kappa, mu, pmix, int.displ)
```

## Arguments

n	number of observations. Ignored if at least one of the other parameters have length $k > 1$ , in which case, all the parameters are recycled to length $k$ to produce $k$ random variates.
kappa	vector of component concentration (inverse-variance) parameters, $\text{kappa} > 0$ .
mu	vector of component means.
pmix	vector of mixing proportions.
x	vector of angles (in radians) where the densities are to be evaluated.
int.displ	integer displacement. The allowed values are 1, 2, 3, 4 and 5. Default is 3.

## Details

pmix, mu and kappa must be of the same length, with  $j$ -th element corresponding to the  $j$ -th component of the mixture distribution.

The univariate wrapped normal mixture distribution with component size  $K = \text{length}(\text{pmix})$  has density

$$g(x) = p[1] * f(x; \kappa[1], \mu[1]) + \dots + p[K] * f(x; \kappa[K], \mu[K])$$

where  $p[j]$ ,  $\kappa[j]$ ,  $\mu[j]$  respectively denote the mixing proportion, concentration parameter and the mean parameter for the  $j$ -th component and  $f(\cdot; \kappa, \mu)$  denotes the density function of the (univariate) wrapped normal distribution with mean parameter  $\mu$  and concentration parameter  $\kappa$ .

## Value

dwnormmix computes the density and rwnormmix generates random deviates from the mixture density.

**Examples**

```

kappa <- 1:3
mu <- 0:2
pmix <- c(0.3, 0.3, 0.4)
x <- 1:10
n <- 10

# mixture densities calculated at each point in x
dwnormmix(x, kappa, mu, pmix)

# number of observations generated from the mixture distribution is n
rwnormmix(n, kappa, mu, pmix)

```

---

summary.angmcmc

*Summary statistics for parameters from an angmcmc object*


---

**Description**

Summary statistics for parameters from an angmcmc object

**Usage**

```

## S3 method for class 'angmcmc'
summary(object, burnin = 1/3, thin = 1, ...)

```

**Arguments**

object	angular MCMC object.
burnin	initial fraction of the MCMC samples to be discarded as burn-in. Must be a value in [0, 1).
thin	positive integer. If thin = $n$ , only every $n$ -th realizations of the Markov chain is kept.
...	additional arguments affecting the summary produced.

**Details**

Computes (after thinning and discarding burn-in) point estimates with 95% posterior credible sets for all components and all parameters, together with the sample averages of log likelihood and log posterior density.

**Value**

Returns a list with elements estimate, lower, upper, llik and lpd.

### Examples

```
# illustration only - more iterations needed for convergence
fit.vmsin.20 <- fit_vmsinmix(tim8, ncomp = 3, n.iter = 20,
                           ncores = 1)
summary(fit.vmsin.20)
```

---

tim8

*Backbone Dihedral Angles of Triose Phosphate Isomerase (8TIM)*

---

### Description

A dataset consisting of 490 pairs of backbone dihedral angles (in radian scale  $[0, 2\pi)$ )  $(\phi, \psi)$  for the protein Triose Phosphate Isomerase (8TIM). The angles were obtained first by using the DSSP software on the PDB file for 8TIM to get the backbone angles (in degrees), and then by converting all angles into radians.

### Usage

```
data(tim8)
```

### Format

A data frame with 490 rows and 2 variables (backbone dihedral angles) phi and psi.

### Source

8TIM PDB file: <http://www.rcsb.org/pdb/explore.do?structureId=8tim>.

DSSP software: <http://swift.cmbi.ru.nl/gv/dssp/>.

---

WAIC

*Watanabe-Akaike Information Criterion (WAIC) for angmcmc objects*

---

### Description

Watanabe-Akaike Information Criterion (WAIC) for angmcmc objects

### Usage

```
WAIC(object, form = 1, burnin = 1/3, thin = 1, ...)
```





---

wind

*Saturna Island wind directions*

---

**Description**

A dataset consisting of 239 observations on wind direction in radians (original measurements were in 10s of degrees), measured at Saturna Island, British Columbia, Canada during October 1-10, 2016 (obtained from Environment Canada website). There was a severe storm during October 4-7, which caused significant fluctuations among the wind directions.

**Usage**

```
data(wind)
```

**Format**

An object of class `numeric` of length 239.

**Source**

Environment Canada: [http://climate.weather.gc.ca/climate\\_data/data\\_quality\\_e.html](http://climate.weather.gc.ca/climate_data/data_quality_e.html).

CBC news on the storm: <http://www.cbc.ca/news/canada/british-columbia/storm-bc-1.3795204>.

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